Multiple Linear Regression: collinearity, model selection

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Today's topics

- collinearity and non-identifiability
- categorical predictors

Example: predicting respiratory disease severity ("lung" dataset)

Multiple linear regression model

Observe data (y, x₁,..., x_p). Want to estimate β₀, β₁,..., β_p in the model

$$\hat{y} = \beta_0 + \beta_1 x_1 + \ldots + \beta_\rho x_\rho$$

Assumptions

- Residuals have mean zero, constant variance, are independent.
- Model is true.

As in simple linear regression, we want to find the β that minimizes the residual sum of squares.

$$RSS(\beta) = \sum_{i} \epsilon_i^2 = \sum_{i} (\hat{y}_i - y)^2$$

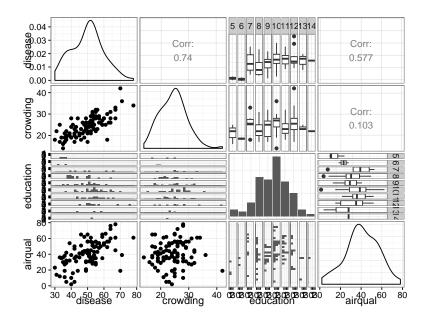
Lung Data Example

99 observations on patients who have sought treatment for the relief of respiratory disease symptoms.

```
dat <- read.table("lungc.txt", header=TRUE)
dat$education <- factor(dat$education)</pre>
```

The variables are:

- disease measure of disease severity (larger values indicates more serious condition).
- education highest grade completed
- crowding measure of crowding of living quarters (larger values indicate more crowding)
- airqual measure of air quality at place of residence (larger number indicates poorer quality)
- nutrition nutritional status (larger number indicates better nutrition)
- smoking smoking status (1 if smoker, 0 if non-smoker)



Lung Data Example

mlr1 <- lm(disease ~ crowding + education + airqual, data=dat)
summary(mlr1)\$coef</pre>

##		Estimate	Std. Error	t value	Pr(> t)
##	(Intercept)	-0.3586956	2.76552352	-0.1297026	8.971011e-01
##	crowding	1.3262116	0.08234084	16.1063641	7.740071e-28
##	education6	-0.7336823	3.35015465	-0.2189995	8.271634e-01
##	education7	2.9068360	2.70190821	1.0758456	2.849708e-01
##	education8	3.1398986	2.34214900	1.3406058	1.835386e-01
##	education9	5.6692646	2.36109920	2.4011124	1.847505e-02
##	education10	5.7785688	2.36728284	2.4410133	1.667257e-02
##	education11	8.2823722	2.43116863	3.4067453	9.972669e-04
##	education12	8.1534760	2.51262476	3.2450034	1.667638e-03
##	education13	13.2612311	2.99374060	4.4296527	2.733931e-05
##	education14	12.8540674	4.23758361	3.0333484	3.187806e-03
##	airqual	0.2950850	0.02549204	11.5755737	2.660791e-19

Least squares estimates: identifiability issues

If two of your variables are identical, or simple transformations of one another, least squares won't work

- This means that there will be an infinite number of mathematically equivalent least squares solutions.
- In practice, true non-identifiability (there really are infinite solutions) is rare.
- Can happen if X is not of full rank, i.e. the columns of X are linearly dependent (for example, including weight in Kg and lb as predictors)
- Can happen if there are fewer data points than terms in X:
 n
- More common, and perhaps more dangerous, is collinearity.

Infinite solutions

Suppose I fit a model $y_i = \beta_0 + \beta_1 x_{i1} + \epsilon_i$.

- I have estimates $\hat{\beta}_0 = 1, \hat{\beta}_1 = 2$
- I put in a new variable $x_2 = x_1$
- My new model is $y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \epsilon_i$

Possible least squares estimates that are equivalent to my first model:

$$\hat{\beta}_{0} = 1, \hat{\beta}_{1} = 2, \hat{\beta}_{2} = 0$$

$$\hat{\beta}_{0} = 1, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 2$$

$$\hat{\beta}_{0} = 1, \hat{\beta}_{1} = 1002, \hat{\beta}_{2} = -1000$$

$$\hat{\beta}_{0} = 0, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{2} = 0, \hat{\beta}_{1} = 0, \hat{\beta}_{2} = 0, \hat{\beta}$$

Non-identifiability example: lung data

```
mlr3 <- lm(disease ~ airqual, data=dat)
summary(mlr3)$coef</pre>
```

 ##
 Estimate Std. Error t value
 Pr(>|t|)

 ## (Intercept)
 35.4444812
 2.23127089
 15.885333
 9.706236e-29

 ## airqual
 0.3537389
 0.05085138
 6.956329
 4.105421e-10

```
dat$x2 <- dat$airqual/100
mlr4 <- lm(disease ~ airqual + x2, data=dat)
summary(mlr4)$coef</pre>
```

 ##
 Estimate Std. Error t value
 Pr(>|t|)

 ## (Intercept)
 35.4444812
 2.23127089
 15.885333
 9.706236e-29

 ## airqual
 0.3537389
 0.05085138
 6.956329
 4.105421e-10

Non-identifiablity: causes and solutions

- Often due to data coding errors (variable duplication, scale changes)
- Pretty easy to detect and resolve
- Can be addressed using *penalties* (might come up much later)
- A bigger problem is near-unidentifiability (collinearity)

Diagnosing collinearity

- Arises when variables are highly correlated, but not exact duplicates
- Commonly arises in data (perfect correlation is usually there by mistake)
- Might exist between several variables, i.e. a linear combination of several variables exists in the data
- A variety of tools exist (correlation analyses, multiple R², eigen decompositions)

Effects of collinearity

Suppose I fit a model $y_i = \beta_0 + \beta_1 x_{i1} + \epsilon_i$.

- I have estimates $\hat{\beta}_0 = 1, \hat{\beta}_1 = 2$
- I put in a new variable $x_2 = x_1 + error$, where *error* is pretty small
- My new model is $y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \epsilon_i$
- Possible least squares estimates that are nearly equivalent to my first model:

$$\hat{\beta}_0 = 1, \hat{\beta}_1 = 2, \hat{\beta}_2 = 0 \hat{\beta}_0 = 1, \hat{\beta}_1 = 0, \hat{\beta}_2 = 2 \hat{\beta}_0 = 1, \hat{\beta}_1 = 1002, \hat{\beta}_2 = -1000 \dots$$

A unique solution exists, but it is hard to find

Effects of collinearity

- Collinearity results in a "flat" RSS
- Makes identifying a unique solution difficult
- Dramatically inflates the variance of LSEs

Collinearity example: lung data

dat\$crowd2 <- dat\$crowding + rnorm(nrow(dat), sd=.1)
mlr5 <- lm(disease ~ crowding + airqual, data=dat)
summary(mlr5)\$coef</pre>

Estimate Std. Error t value Pr(>|t|)
(Intercept) 2.8841197 2.49069149 1.157959 2.497533e-01
crowding 1.4027587 0.09341356 15.016650 6.154176e-27
airqual 0.3104388 0.02808020 11.055436 8.202723e-19

mlr6 <- lm(disease ~ crowding + crowd2 + airqual, data=dat)
summary(mlr6)\$coef</pre>

 ##
 Estimate Std. Error
 t value
 Pr(>|t|)

 ## (Intercept)
 2.8737453
 2.5039798
 1.1476711
 2.539863e-01

 ## crowding
 0.6121638
 4.3473948
 0.1408117
 8.883169e-01

 ## crowd2
 0.7918196
 4.3531139
 0.1818973
 8.560509e-01

 ## airqual
 0.3101664
 0.0282624
 10.9745229
 1.391889e-18

Using Variance Inflation Factors: lung data

VIFs find variables that are highly related.

The VIF for the k^{th} predictor in your model is

$$VIF_k = rac{1}{1-R_k^2}$$

where R_k^2 is the R^2 from the model with X_k as the response and all other X variables as the predictors.

```
car::vif(mlr5)
## crowding airqual
## 1.010657 1.010657
car::vif(mlr6)
## crowding crowd2 airqual
## 2166.934175 2167.436307 1.013503
```

Rule of thumb is that if any $VIF_k \downarrow 10$, then you should be

Some take away messages

- Collinearity can (and does) happen, so be careful
- Often contributes to the problem of variable selection, which we'll touch on later

Model selection

Why are you building a model in the first place?

Model selection: considerations

Things to keep in mind...

- Why am I building a model? Some common answers
 - Estimate an association
 - Test a particular hypothesis
 - Predict new values
- What predictors will I allow?
- What predictors are needed?

Different answers to these questions will yield different final models.

Model selection: realities

All models are wrong. Some are more useful than others. - George Box

- In practice, issues with sample size, collinearity, and available predictors are real problems.
- There is not a single best algorithm for model selection! It pretty much always requires thoughful reasoning and knowledge about the data at hand.
- When in doubt (unless you are specifically "data mining"), err on the side creating a process that does not require choices being made (by you or the computer) about which covariates to include.

Basic ideas for model selection

For association studies, when your sample size is large

- Include key covariates of interest.
- Include covariates needed because they might be confounders.
- Include covariates that your colleagues/reviewers/collaborators will demand be included for face validity.
- Do NOT go on a fishing expedition for significant results!
- Do NOT use "stepwise selection" methods!
- Subject the selected model to model checking/diagnostics, possibly adjust model structure (i.e. include non-linear relationships with covariates) as needed.

Basic ideas for model selection

For association studies, when your sample size is small

- Same as above, but may need to be more frugal with how many predictors you include.
- Rule of thumb for multiple linear regression is to have at least 15 observations for each covariate you include in your model.

Today's big ideas

- dangers of collinearity and non-identifiability
- model selection

Analyze the NHANES dataset. Create a parsimonious model with the outcome variable of cholesterol (chol) that estimates relationships with other variables in the dataset. Justify your choices of which covariates you included using some basic knowledge about what factors might impact cholesterol levels.

library(NHANES)
data(NHANES)
?NHANES