Simple Linear Regression and the Method of Least Squares

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This material is part of the **statsTeachR** *project*

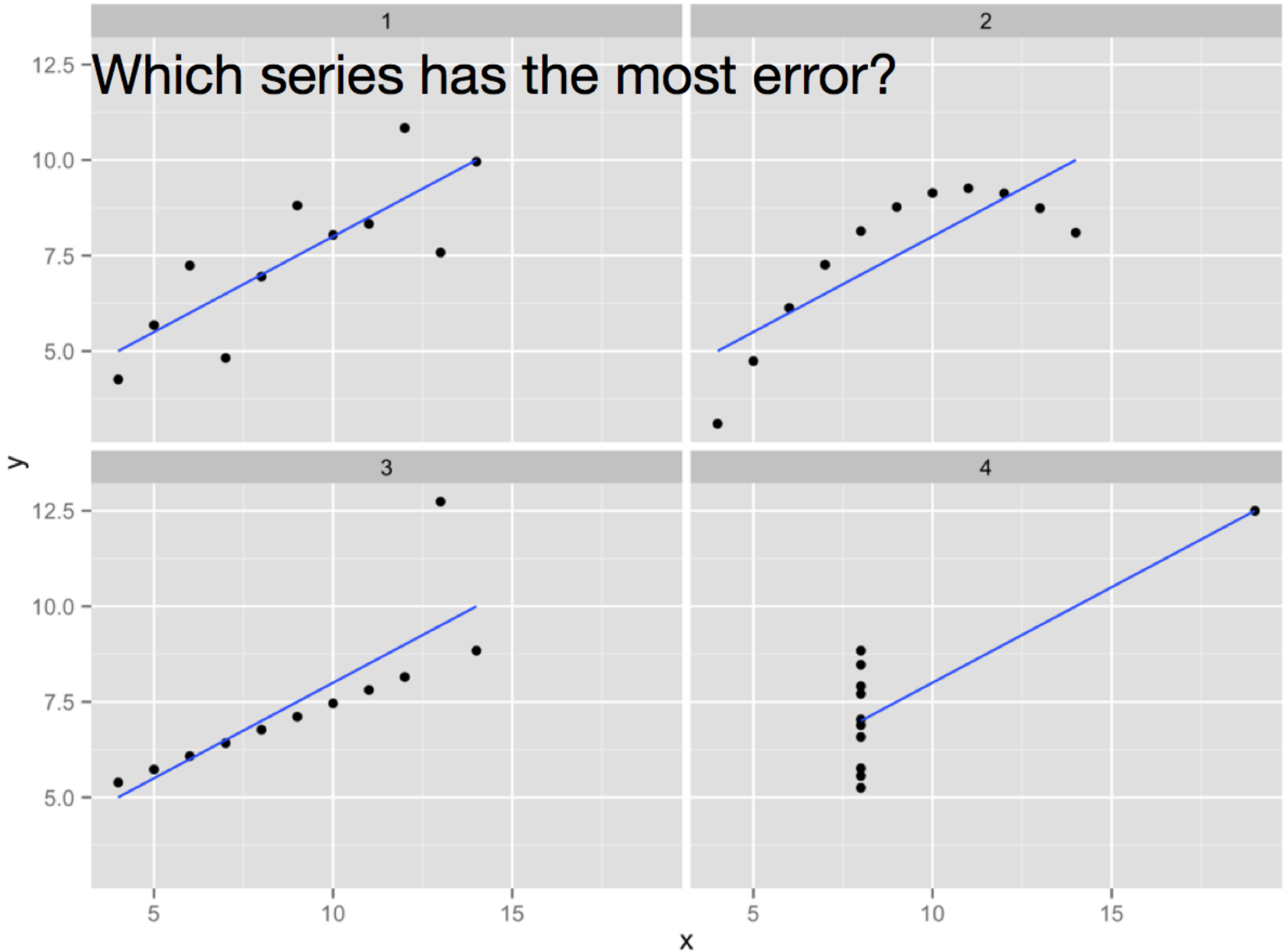
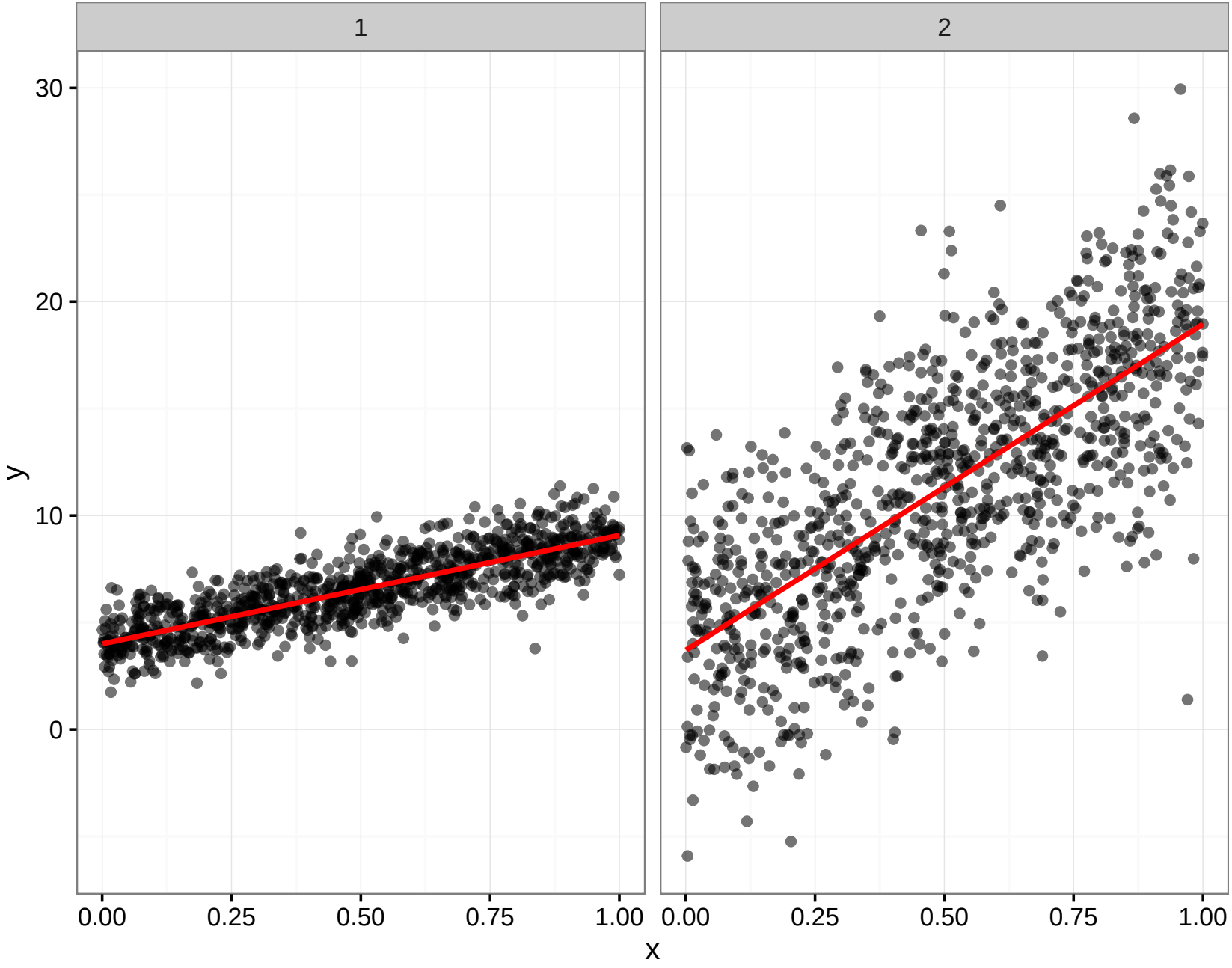


Figure acknowledgements to Hadley Wickham.

Which data show a stronger association?



Goals for this class

You should be able to...

- interpret regression coefficients.
- derive estimators for SLR coefficients.
- implement a SLR from scratch (i.e. not using `lm()`).
- explain why some points have more influence than others on the fitted line.

Regression modeling

- Want to use predictors to learn about the outcome distribution, particularly conditional expected value.
- Formulate the problem parametrically

$$\mathbb{E}(y \mid x) = f(x; \beta) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots$$

- (Note that other useful quantities, like covariance and correlation, tell you about the joint distribution of y and x)

Brief Detour: Covariance and Correlation

$$
\begin{aligned}
cov(x, y) &= \mathbb{E}\left[(x - \mu_x)(y - \mu_y)\right] \\
cor(x, y) &= \frac{cov(x, y)}{\sqrt{var(x)var(y)}}
\end{aligned}
$$

Simple linear regression

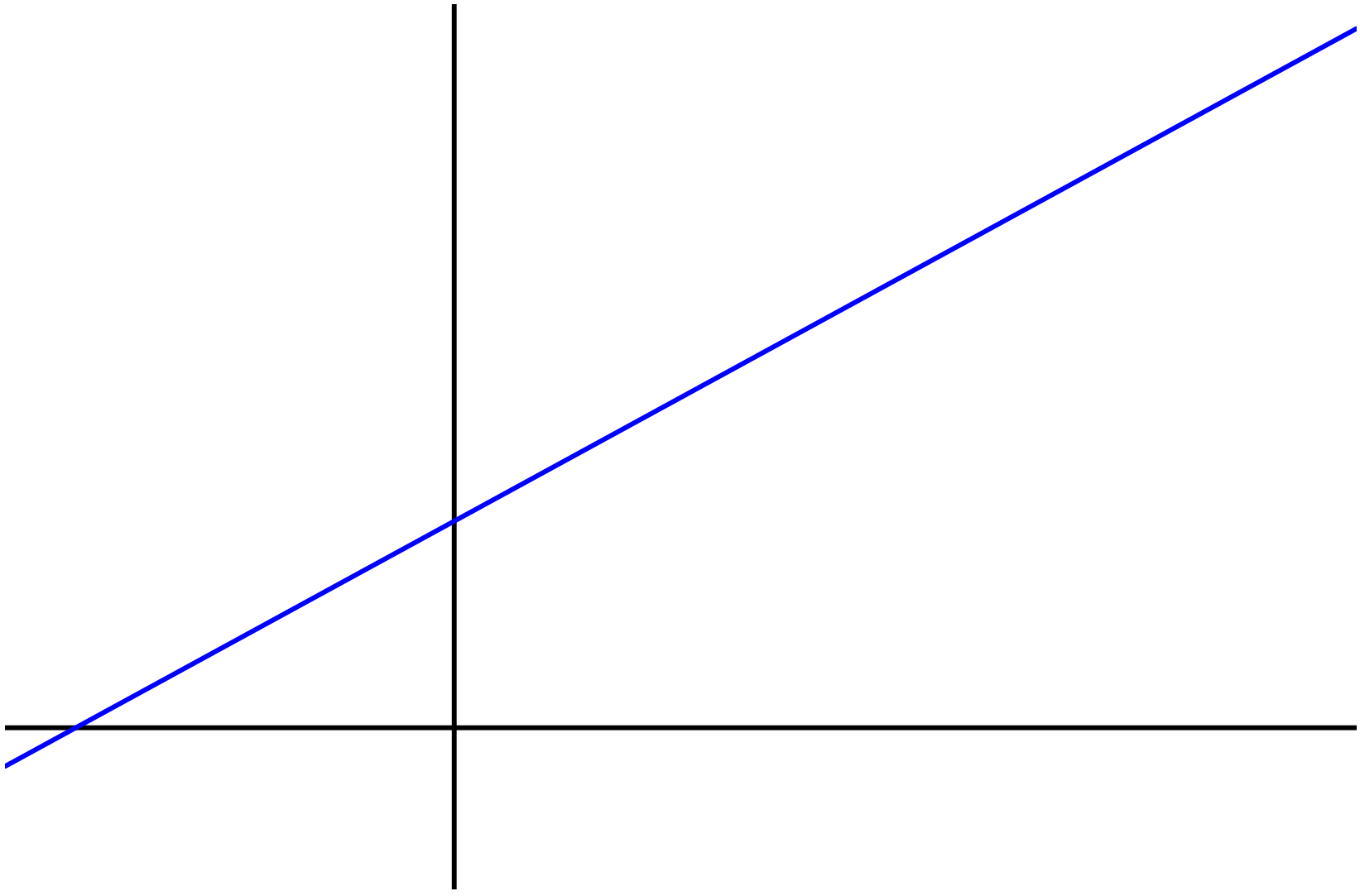
- Linear models are a special case of all regression models; simple linear regression is the simplest place to start
- Only one predictor:

$$\mathbb{E}(y \mid x) = f(x; \beta) = \beta_0 + \beta_1 x_1$$

- Useful to note that $x_0 = 1$ (implicit definition)
- Somehow, estimate β_0, β_1 using observed data.

Coefficient interpretation

Coefficient interpretation



Step 1: Always look at the data!

- Plot the data using, e.g. the `plot()` or `qplot()` functions
- Do the data look like the assumed model?
- Should you be concerned about outliers?
- Define what you expect to see before fitting any model.

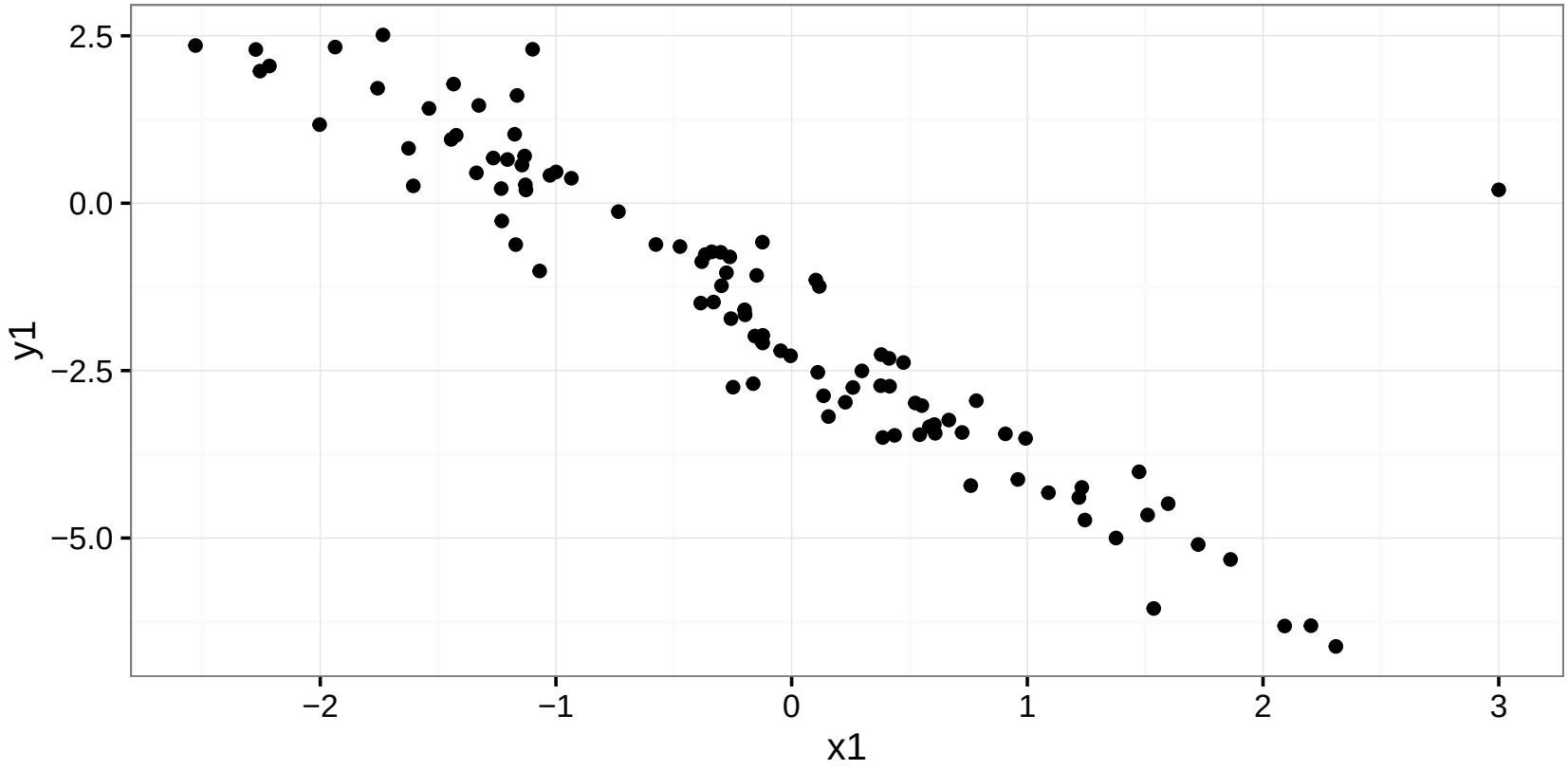
Least squares estimation

- Observe data (y_i, x_i) for subjects $1, \ldots, I$. Want to estimate β_0, β_1 in the model

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i; \ \epsilon_i \overset{iid}{\sim} (0, \sigma^2)$$

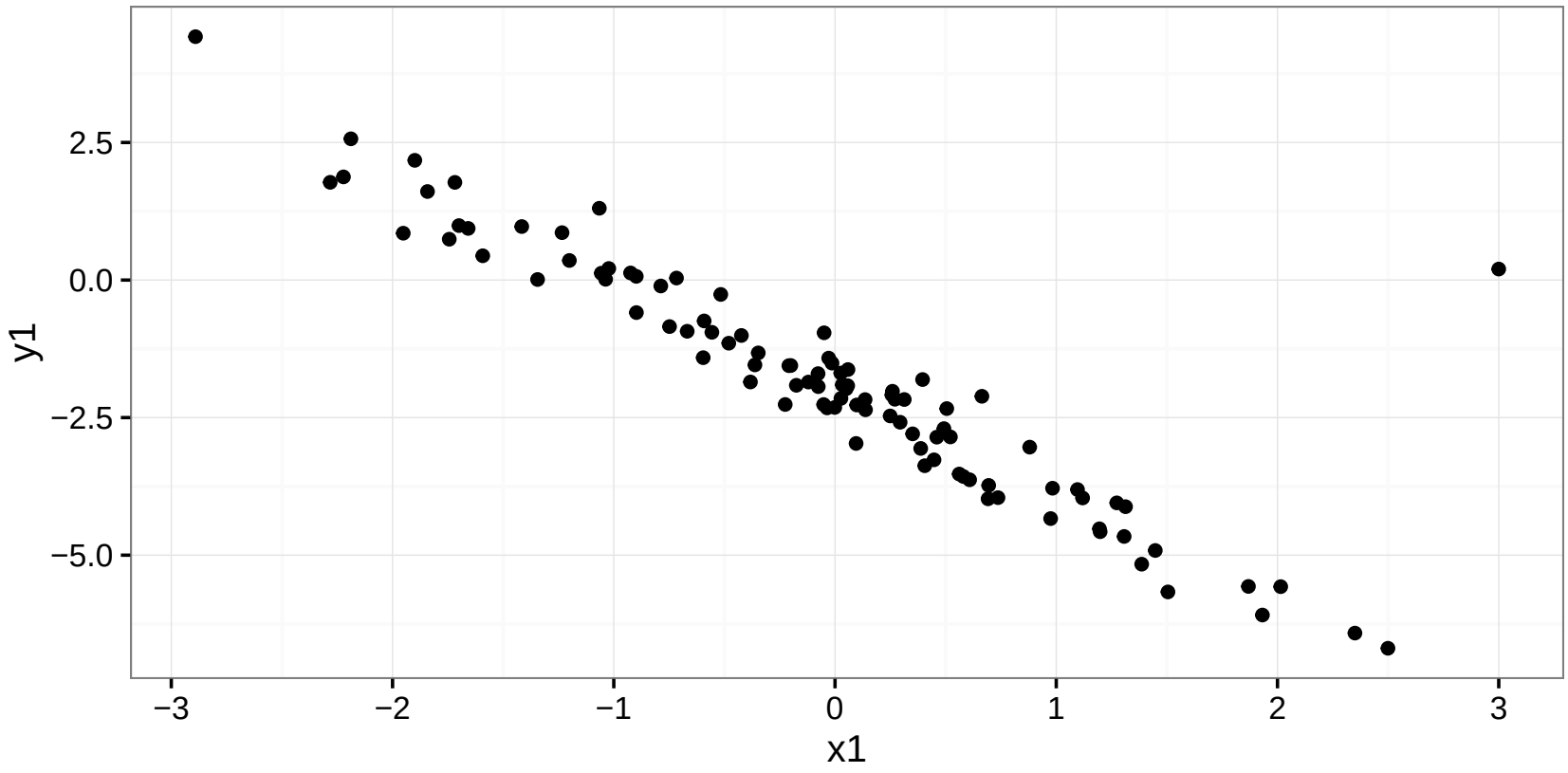
- Recall the assumptions:
  - A1: The model: e.g. $y_i = f(x_i; \beta) + \epsilon_i = \beta_0 + \beta_1 x_{i,1} + \epsilon_i$
  - A2: Unbiased errors: $\mathbb{E}[\epsilon_i | x_i] = \mathbb{E}[\epsilon_i] = 0$
  - A3: Uncorrelated errors: $cov(\epsilon_i, \epsilon_j) = 0$ for $i \neq j$.
  - A4: Constant variance: $Var[y_i | x_i] = \sigma^2$
  - A5: Probability distribution: e.g. $\epsilon_i \overset{iid}{\sim} N(0, \sigma^2)$ [not needed for LS, is needed for inference].
  - A6: Representative sampling: generalize to population.

Any violations of assumptions? (Ex. 1)

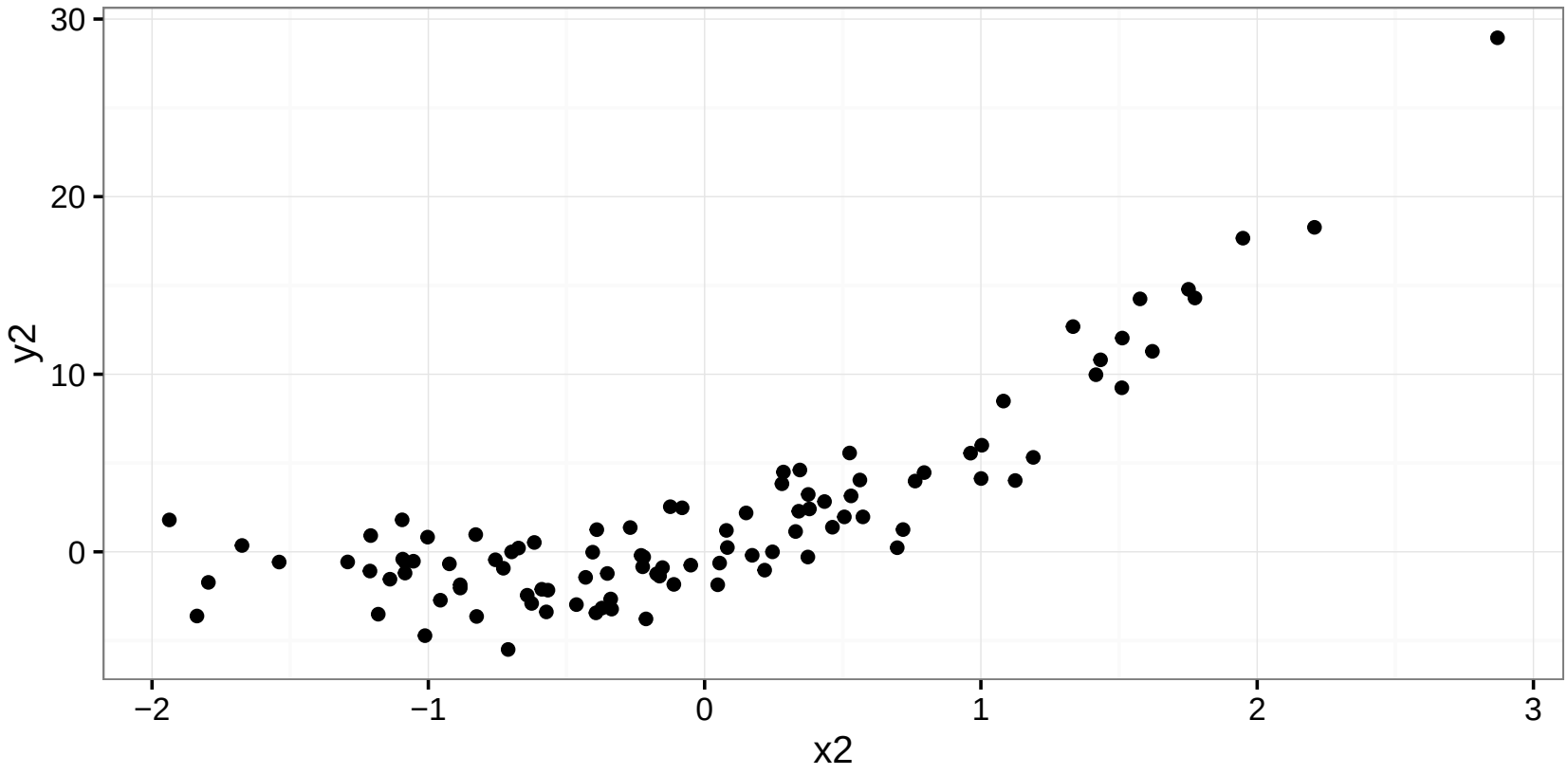


Any violations of assumptions? (Ex. 1)

```
x1 <- rnorm(100)
y1 = -2-2*x1 + rnorm(100, 0, .5)
x1[1] <- 3; y1[1] <- .2
qplot(x1, y1)
```

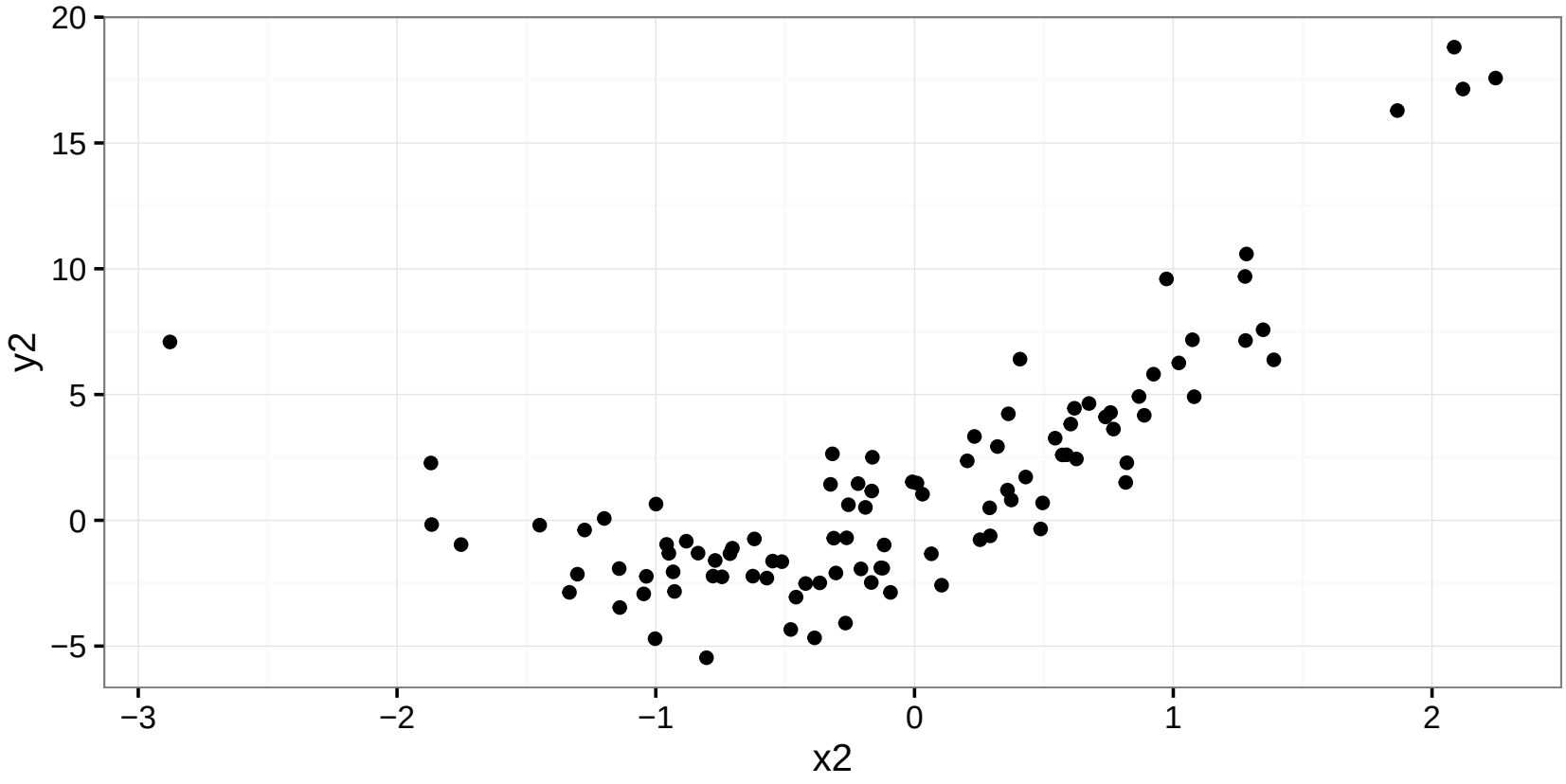


Any violations of assumptions? (Ex. 2)

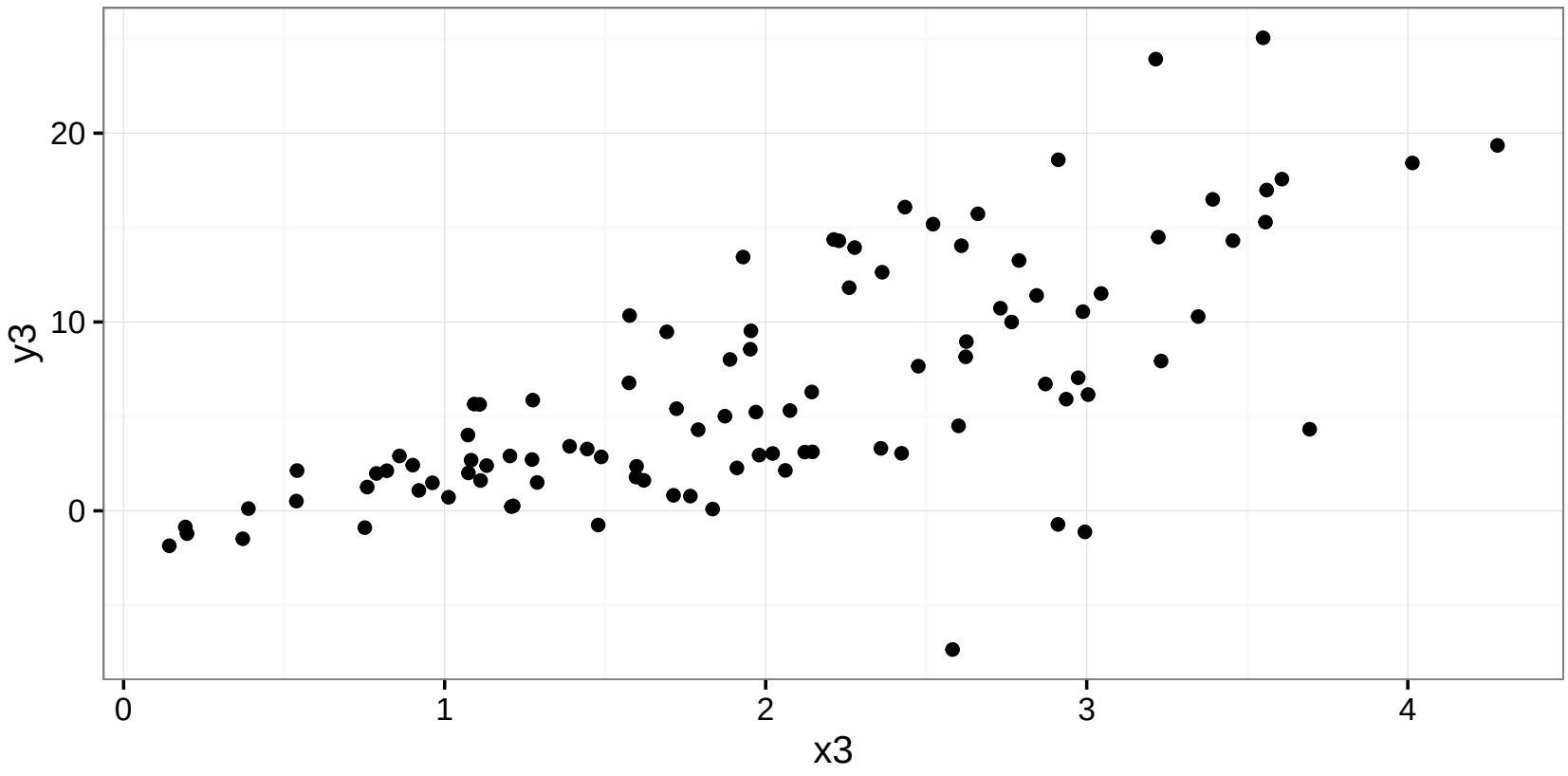


Any violations of assumptions? (Ex. 2)

```
x2 <- rnorm(100)
y2 = -2+2*(x2+1)^2 + rnorm(100, 0, 2)
qplot(x2, y2)
```

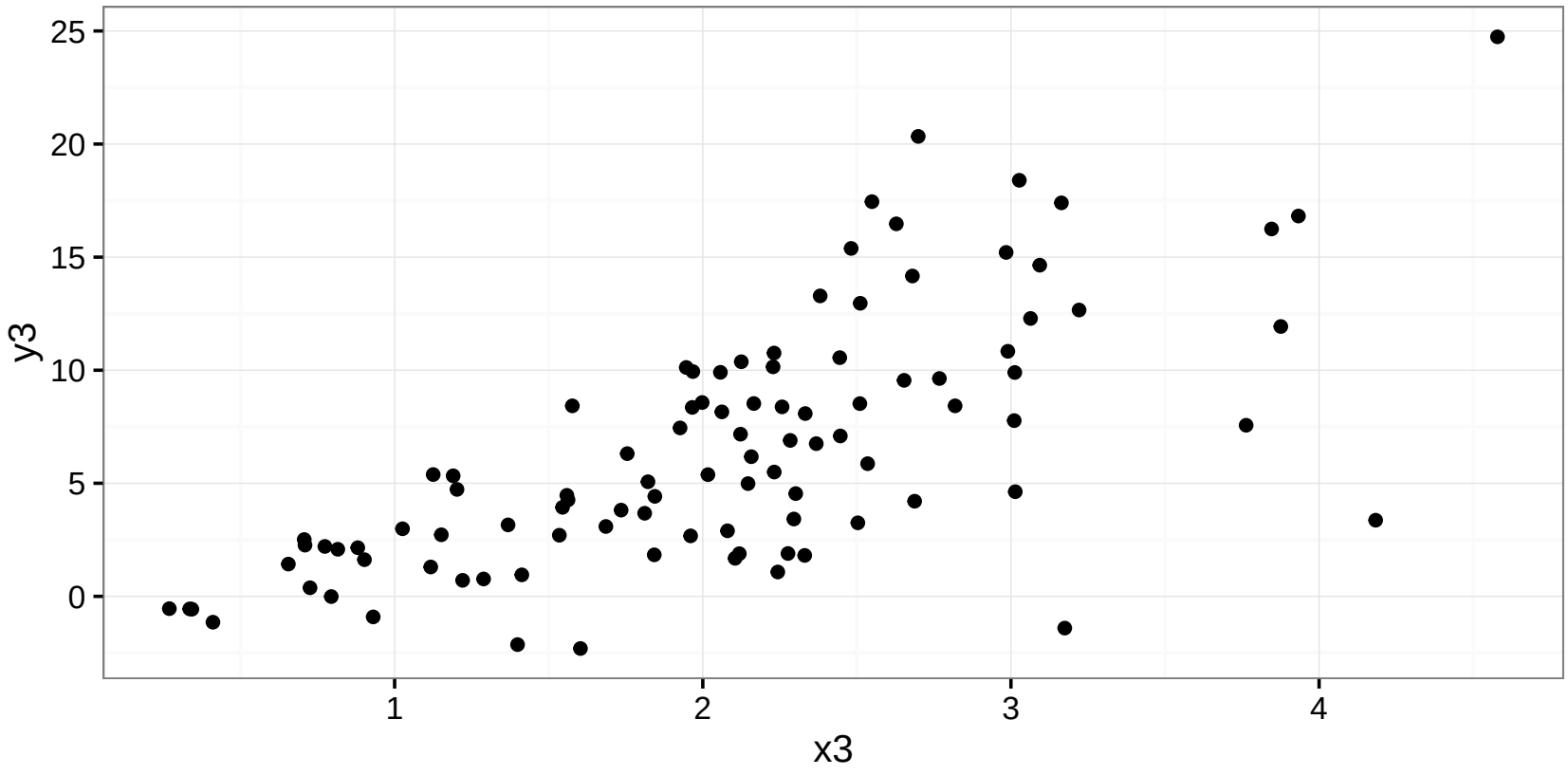


Any violations of assumptions? (Ex. 3)



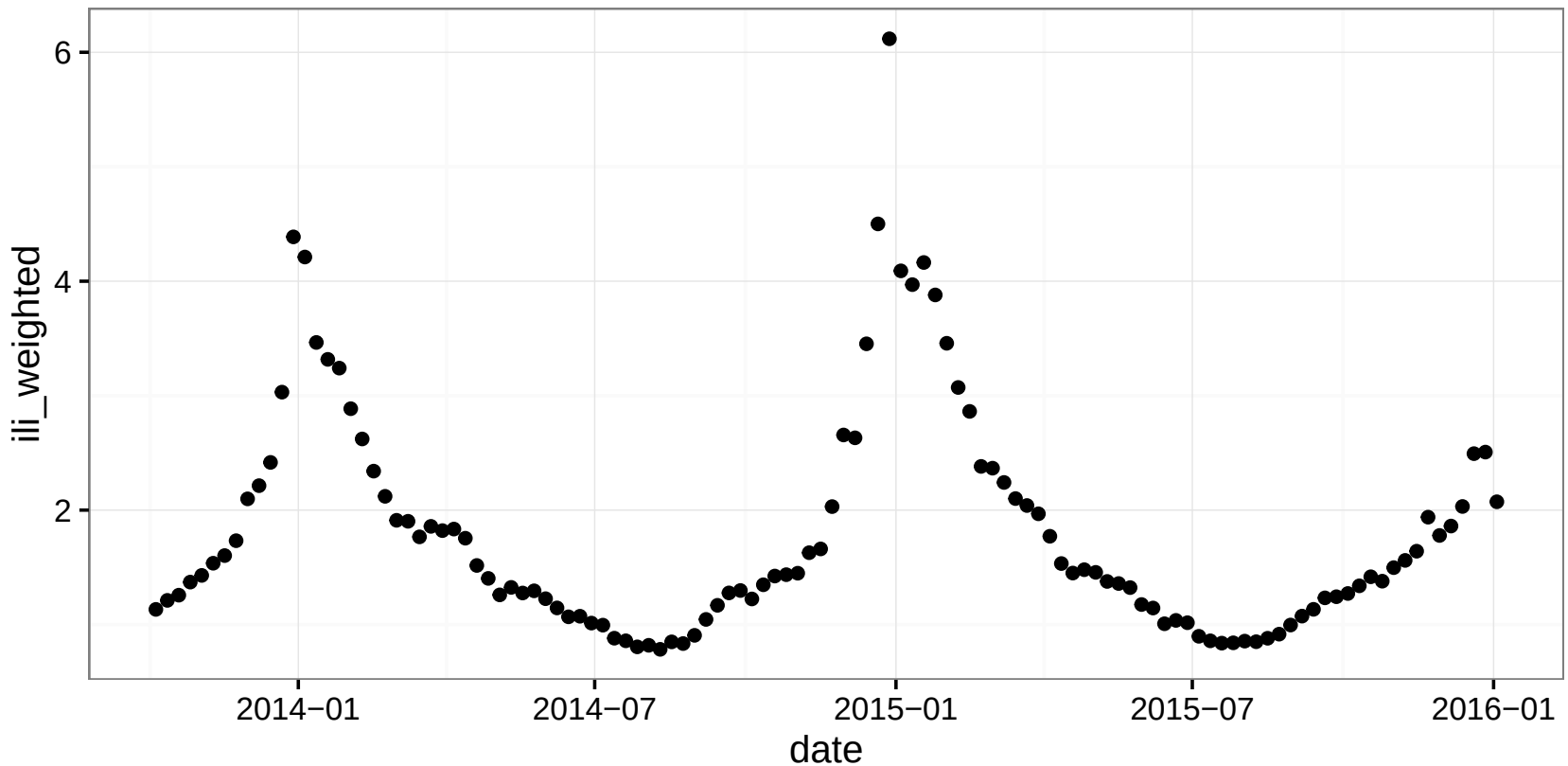
Any violations of assumptions? (Ex. 3)

```
x3 <- abs(rnorm(100, mean=2))
y3 = -2+4*x3 + rnorm(100, 0, x3*2)
qplot(x3, y3)
```



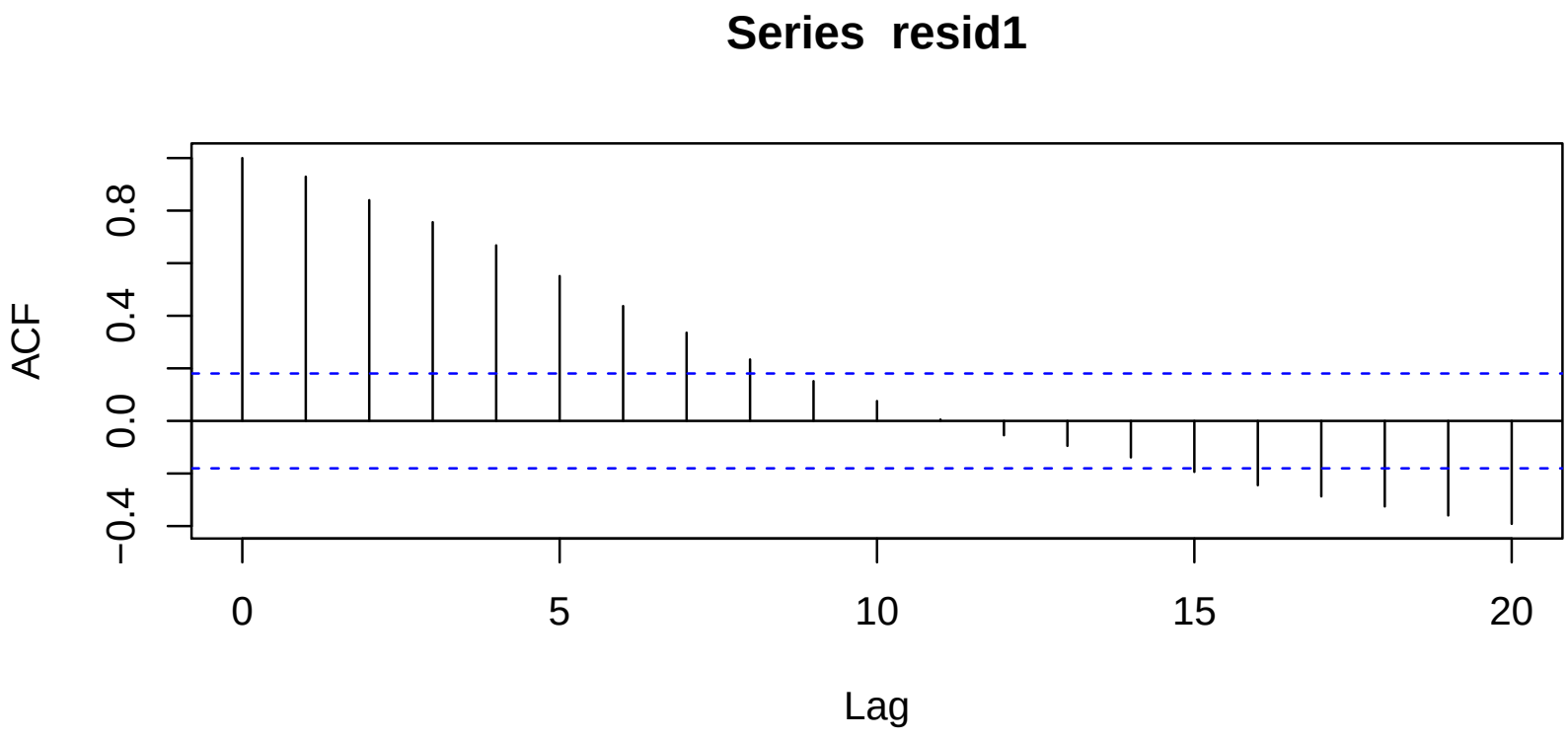
Any violations of assumptions? (Ex. 4)

```
library(cdcfluview)
library(dplyr)
usflu <- get_flu_data("national", "ilinet", years=2013:2015)
usflu <- mutate(usflu,
                date = as.Date(paste0(YEAR, sprintf("%02d", WEEK), "00"),
                               format="%Y%W%w"),
                ili_weighted = X.UNWEIGHTED.ILI)
ggplot(usflu, aes(x=date, y=ili_weighted)) + geom_point()
```

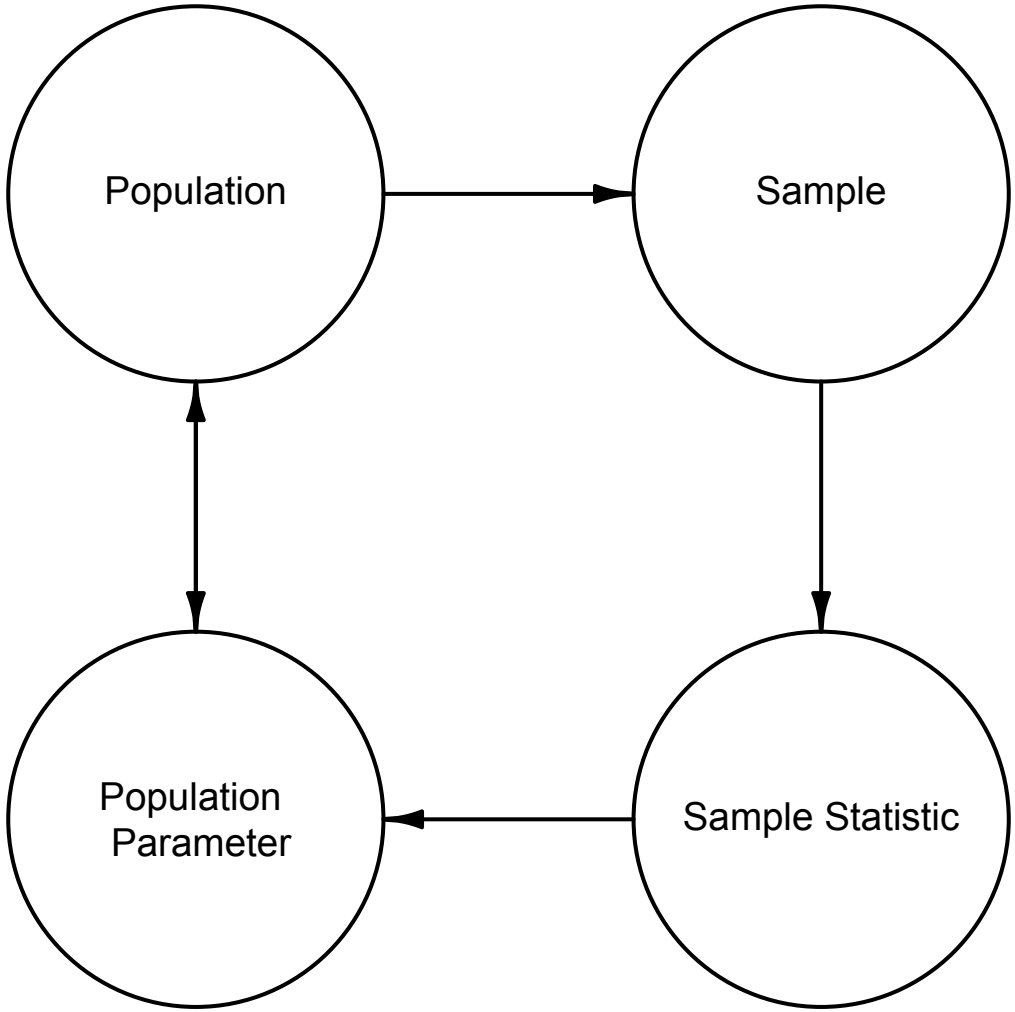


Any violations of assumptions? (Ex. 4)

```
fm1 <- lm(ili_weighted~date, data=usflu)
resid1 <- resid(fm1)
acf(resid1)
```



Circle of Life



Least squares estimation

- Recall that for a single sample $y_i, i \in 1, \ldots, N$, the sample mean $\hat{\mu}_y$ minimizes the sum of squared deviations.

$$RSS(\mu_y) = \sum_{i=1}^{N} (y_i - \mu_y)^2$$

Least squares estimation

Find $\hat{\beta}_0$ and β_1. By minimizing RSS relative to each parameter.

$$RSS(\beta_0, \beta_1) = \sum_{i=1}^{N} (y_i - \mathbb{E}[y_i|x_i])^2$$

We obtain

$$\begin{aligned} \hat{\beta}_0 &= b_0 = \bar{y} - b_1 \bar{x} \\ \hat{\beta}_1 &= b_1 = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2} \end{aligned}$$

Notes about LSE

Relationship between correlation and slope

$$\rho = \frac{cov(x,y)}{\sqrt{var(x)var(y)}}; \qquad \beta_1 = \frac{cov(x,y)}{var(x)}$$

Why we need to keep watch for outliers

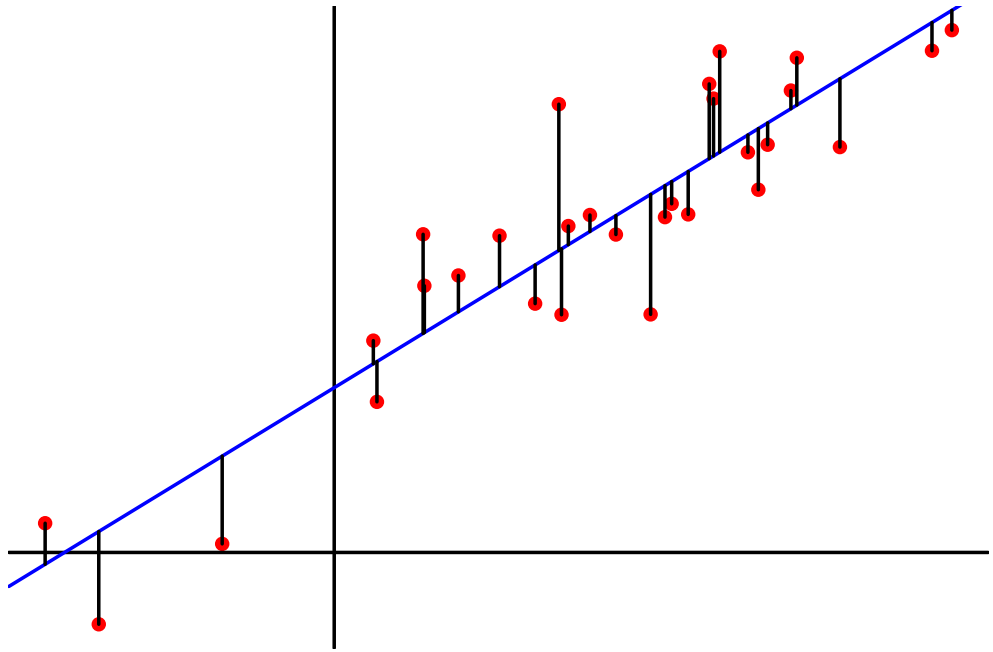
$$\begin{aligned}
\hat{\beta}_1 &= \frac{\sum(y_i - \bar{y})(x_i - \bar{x})}{\sum(x_i - \bar{x})^2} \\
&= \frac{\sum \frac{y_i - \bar{y}}{x_i - \bar{x}}(x_i - \bar{x})^2}{\sum(x_i - \bar{x})^2} \\
&= \sum \frac{y_i - \bar{y}}{x_i - \bar{x}}\omega_i
\end{aligned}$$

Note that weight ω_i increases as x_i gets further away from $\bar{x}$.

Geometric interpretation of least squares

Least squares minimizes the sum of squared vertical distances between observed and estimated y's:

$$\underset{\beta_0, \beta_1}{min} \sum_{i=1}^{I} (y_i - (\beta_0 + \beta_1 x_i))^2$$



Least squares foreshadowing

- Didn't have to choose to minimize squares – could minimize absolute value, for instance.
- Least squares estimates turn out to be a "good idea" – unbiased, BLUE (Best Linear Unbiased Estimator).
- Later we'll see about maximum likelihood as well.

Lab exercise: computing $\hat{\beta}$ on your own

- Load the heights data from lecture 1.
- Run a linear model using the R function `lm()`, with daughter height as the outcome.
- Compare the results of that regression with hand-calculated $\hat{\beta}_0$ and $\hat{\beta}_1$ coefficients.

```
# sample code
install.packages("alr3")
library(alr3)
data(heights)
fm1 <- lm(Dheight ~ Mheight, data=heights)
summary(fm1)
```